

Problem Statement

We consider **sequential Bayesian experimental design (BED)** for partially observable state-space models (SSMs) with a *design* input ξ_t that shapes data acquisition:

$$\text{(transition)} \quad \mathbf{x}_t \sim f(\mathbf{x}_t | \mathbf{x}_{t-1}, \boldsymbol{\theta}, \xi_t), \quad (1)$$

$$\text{(observation)} \quad \mathbf{y}_t \sim g(\mathbf{y}_t | \mathbf{x}_t, \boldsymbol{\theta}, \xi_t). \quad (2)$$

- $\boldsymbol{\theta}$: static parameters – **to learn**
- $\mathbf{x}_t$: latent states – **to learn** (partially observed)
- ξ_t : design variables – **to optimize** (online)

Optimization objective: *expected information gain* (EIG). We choose ξ_t to maximize the expected reduction in uncertainty about $\boldsymbol{\theta}$:

$$\text{EIG}_{\boldsymbol{\theta}}(\xi_t) = \mathbb{E}_{p(\mathbf{y}_t | \xi_t, h_{t-1})} \left[\mathcal{H}[p(\boldsymbol{\theta} | h_{t-1})] - \mathcal{H}[p(\boldsymbol{\theta} | \mathbf{y}_t, \xi_t, h_{t-1})] \right] \quad (3)$$

$$= \mathbb{E}_{p(\boldsymbol{\theta} | h_{t-1}) p(\mathbf{y}_t | \boldsymbol{\theta}, \xi_t)} \left[\log \underbrace{p(\mathbf{y}_t | \boldsymbol{\theta}, \xi_t)}_{\text{likelihood}} - \log \underbrace{p(\mathbf{y}_t | \xi_t)}_{\text{evidence}} \right]. \quad (4)$$

- $\mathcal{H}[\cdot]$ is the Shannon entropy,
- $h_{t-1} = \{\xi_{1:t-1}, \mathbf{y}_{1:t-1}\}$ is the history up to time $t-1$,

→ The **optimal design** ξ_t^* is the one that maximizes $\text{EIG}_{\boldsymbol{\theta}}(\xi_t)$:

$$\xi_t^* = \arg \max_{\xi_t \in \Omega} \text{EIG}_{\boldsymbol{\theta}}(\xi_t) \quad (5)$$

Method: Online BED approach

Challenges:

(i) Intractability and latent-state marginalizations:

$$\begin{aligned} \text{(likelihood)} \quad p(\mathbf{y}_t | \boldsymbol{\theta}, \xi_t) &= \mathbb{E}_{p(\mathbf{x}_{0:t} | \boldsymbol{\theta}, h_{t-1})} [g(\mathbf{y}_t | \mathbf{x}_t, \boldsymbol{\theta}, \xi_t)], \\ \text{(evidence)} \quad p(\mathbf{y}_t | \xi_t) &= \mathbb{E}_{p(\boldsymbol{\theta} | h_{t-1}) p(\mathbf{x}_{0:t} | \boldsymbol{\theta}, h_{t-1})} [g(\mathbf{y}_t | \mathbf{x}_t, \boldsymbol{\theta}, \xi_t)]. \end{aligned}$$

(ii) Sequential inference bottleneck: $p(\mathbf{x}_{0:t}, \boldsymbol{\theta} | h_t)$ changes every step; naïve recomputation would replay the entire history each time.

Key idea:

Extend BED to partial observability by:

- deriving **new Monte Carlo estimators** of $\text{EIG}_{\boldsymbol{\theta}}$ (and its gradient) that treat the latent-state integrals explicitly; and
- leveraging nested particle filters (NPFs) (Crisan and Míguez, 2018), that is **online Bayesian inference methods**, to *reuse* state-parameter particles and *avoid replaying past data*.

TL;DR

- **Problem:** identifying the system parameters $\boldsymbol{\theta}$ *online*, where latent states $\mathbf{x}_t$ evolve over time, observed via noisy measurements $\mathbf{y}_t$. These measurement (data acquisition) processes can be influenced by *designs* ξ_t .
- **Proposed solution:** choose designs ξ_t by maximizing the *expected information gain* (EIG) about $\boldsymbol{\theta}$, marginalizing out latent states $\mathbf{x}_t$. Leverage a nested particle filters (NPF) to approximate the intractable posterior $p(\mathbf{x}_t, \boldsymbol{\theta} | h_t)$ *online*.

Algorithm Design selection at time t

- 1: **Input:** samples $\{\boldsymbol{\theta}_{t-1}^{(m)}, \mathbf{x}_{t-1}^{(m,n)}\}_{m=1, n=1}^{M, N} \sim p(\mathbf{x}_{t-1}, \boldsymbol{\theta} | h_{t-1})$;
- 2: initial design $\xi_t^{(0)}$; stepsizes $\{\eta_k\}_{k=0}^{K-1}$
- 3: **Output:** samples $\{\boldsymbol{\theta}_t^{(m)}, \mathbf{x}_t^{(m,n)}\}_{m=1, n=1}^{M, N} \sim p(\mathbf{x}_t, \boldsymbol{\theta} | h_t)$; design ξ_t^*
- 4:
- 5: **for** $k = 0, \dots, K-1$ **do** ▷ design optimization loop
- 6: **(1)** Draw $(\tilde{\mathbf{y}}_t^{(\ell)}, \tilde{\mathbf{x}}_t^{(\ell)}, \boldsymbol{\theta}_{t-1}^{(\ell)})$ from
- 7: $g(\cdot | \mathbf{x}_{t-1}, \boldsymbol{\theta}_{t-1}, \xi_t^{(k)}) f(\cdot | \mathbf{x}_{t-1}, \boldsymbol{\theta}_{t-1}, \xi_t^{(k)}) p(\mathbf{x}_{0:t-1}, \boldsymbol{\theta} | h_{t-1})$
- 8: **(2)** Approximate $p(\tilde{\mathbf{y}}_t | \boldsymbol{\theta}, \xi_t^{(k)})$, $p(\tilde{\mathbf{y}}_t | \xi_t^{(k)})$, and $\widehat{\nabla}_{\xi_t} \text{EIG}_{\boldsymbol{\theta}}(\xi_t^{(k)})$
- 9: **(3)** $\xi_t^{(k+1)} \leftarrow (\xi_t^{(k)} + \eta_k \widehat{\nabla}_{\xi_t} \text{EIG}_{\boldsymbol{\theta}}(\xi_t^{(k)}))$
- 10: **end for**
- 11: Set $\xi_t^* \leftarrow \xi_t^{(K)}$; collect $\mathbf{y}_t$ ▷ observe new data
- 12:
- 13: Jittering: $\boldsymbol{\theta}_{t|t-1}^{(m)} \sim \kappa_M(\cdot | \boldsymbol{\theta}_{t-1}^{(m)})$ ▷ update beliefs (NPF step)
- 14: **for** $m = 1$ to M **do**
- 15: Propagate states: $\mathbf{x}_{t|t-1}^{(m,n)} \sim f(\cdot | \mathbf{x}_{t-1}^{(m,n)}, \boldsymbol{\theta}_{t|t-1}^{(m)}, \xi_t^*)$
- 16: Update state weights $w_{\mathbf{x},t}^{(m,n)}$ and resample to obtain $\mathbf{x}_t^{(m,n)}$
- 17: **end for**
- 18: Update parameter weights $w_{\boldsymbol{\theta},t}^{(m)} = w_{\boldsymbol{\theta},t-1}^{(m)} \sum_{n=1}^N w_{\mathbf{x},t}^{(m,n)}$
- 19: Resample to obtain $\boldsymbol{\theta}_t^{(m)}$

Consistency of the EIG estimator

Theorem 1. Let $\widehat{\mathcal{I}}(\xi_t)$ denote the NMC estimator of the EIG, constructed using M parameter particles, N state particles per parameter, and L pseudo-observations. Under mild assumptions, for any t and ξ_t

$$\widehat{\mathcal{I}}(\xi_t) \xrightarrow[L, M, N \rightarrow \infty]{\text{a.s.}} \mathcal{I}(\xi_t).$$

Assumptions: small/regular jittering kernels, Lipschitz state posterior in $\boldsymbol{\theta}$, bounded positive likelihood, integrand of EIG is Lipschitz and square-integrable.

Example: Moving source location

State. A single source moves in the plane such that

$$\mathbf{x}_t = \begin{bmatrix} p_{x,t} \\ p_{y,t} \\ \phi_t \end{bmatrix} = \begin{bmatrix} p_{x,t-1} + v_x \cos(\phi_{t-1}) \\ p_{y,t-1} + v_y \sin(\phi_{t-1}) \\ \phi_{t-1} \end{bmatrix} + \mathbf{w}_t$$

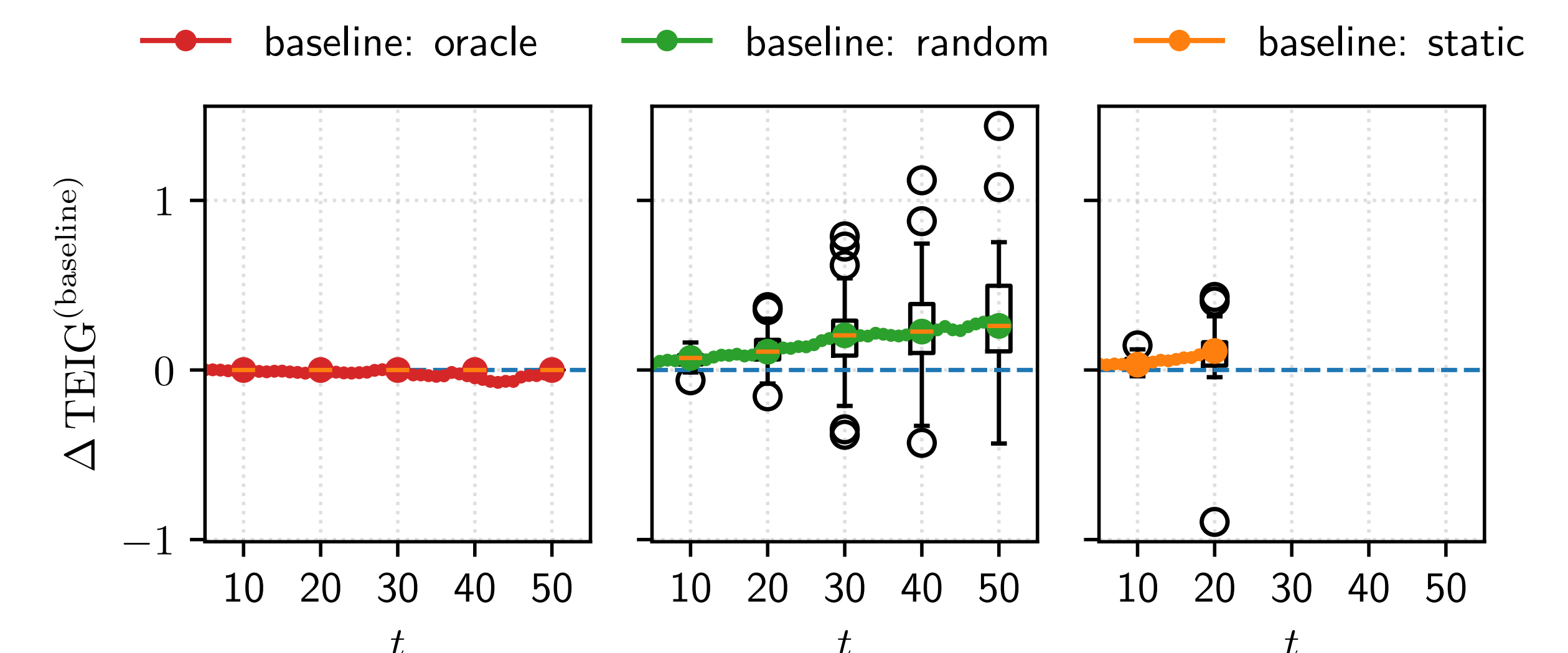
where $\mathbf{w}_t \sim \mathcal{N}(\mathbf{0}, Q)$ and $\boldsymbol{\theta} = (v_x, v_y)$ are unknown parameters.

Observation. Sensors fixed at positions $\{\mathbf{s}_j\}_{j=1}^J \subset \mathbb{R}^2$ return a noisy log-intensity with distance attenuation and cardioid directivity,

$$\log y_{t,j} | \mathbf{x}_t, \xi_t \sim \mathcal{N} \left(\log \left[b + \frac{\alpha_j}{m + \|\mathbf{p}_t - \mathbf{s}_j\|^2} \left(\frac{1 + \cos \Delta_{t,j}(\xi_{t,j})}{2} \right) \right], \sigma^2 \right),$$

$$\Delta_{t,j}(\xi_{t,j}) = \xi_{t,j} - \text{atan2}((\mathbf{p}_t - \mathbf{s}_j)_y, (\mathbf{p}_t - \mathbf{s}_j)_x).$$

Design. The *design* $\xi_t = (\xi_{t,1}, \dots, \xi_{t,J})$ sets sensor *orientations*.



Assessment. We report *relative improvements*,

$$\Delta \text{TEIG}^{(\text{baseline})} = \sum_{s=1}^t \left(\widehat{\text{EIG}}_{\boldsymbol{\theta}}(\xi_s^{\text{BAD-PODS}}) - \widehat{\text{EIG}}_{\boldsymbol{\theta}}(\xi_s^{\text{baseline}}) \right),$$

against two baselines to highlight adaptive gains:

- *Random* baseline: random designs sampled uniformly from Ω .
- *Static* baseline: a non-adaptive version of our method where the full sequence $\xi_{1:T}$ is optimized offline.
- *Oracle* baseline: evaluates EIG on a dense discretisation of the space (instead of gradient-based optimisation).

References

- [1] D. Crisan and J. Míguez, “Nested particle filters for online parameter estimation in discrete-time state-space Markov models,” *Bernoulli*, vol. 24, no. 4A, pp. 3039–3086, 2018.
- [2] S. Pérez-Vieites, S. Iqbal, S. Särkkä, and D. Baumann, “Online Bayesian experimental design for partially observed dynamical systems,” *arXiv preprint arXiv:2511.04403*, 2025.